

# Essential Mathematics for Global Leaders I

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## Statistics

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Where are we ? Today's plan

# PART II: Statistical inference (推計統計学)

## 4. Null Hypothesis Significant Test (NHST) 帰無仮説検定

4.1 Concepts and 1<sup>st</sup> example: z-test

4.2 chi2 and sample variance (カイ二乗と標本分散)

4.3 the Student t-test (Student t-検定)

4.4 Two-sample t-test (=variance) 対応のないt検定

4.5 Paired difference sampling 対応のあるデータ

4.6 Comparing two population variances: F-test

4.7 chi2-test (goodness-of-fit) カイ二乗(簡単な適合度検定)

4.8 chi2-test of independence 独立性のカイ二乗検定

4.9 One-way ANOVA (F-test) 一元配置分散分析 (F検定)

# Chapter 4: NHST

## Section 4.2 chi2 and sample variance

### $\chi^2$ distribution

- $X_1, \dots, X_k$  i.i.d.r.v. from the standard normal  $N(0,1)$ .

**Theorem 1:** The random variable  $Q_k = X_1^2 + X_2^2 + \dots + X_k^2$  follows a  $\chi^2$ -distribution with  $k$  degrees of freedom (自由度)

- For any  $k \geq 1$ ,  $\int_0^\infty x^{k/2-1} e^{-x/2} dx = \Gamma(k/2) 2^{k/2}$ .

• **Definition:**  $\chi_k^2$  has for pdf:

- $f(x; k) = \frac{x^{k/2-1} e^{-x/2}}{\Gamma(\frac{k}{2}) 2^{k/2}}$  if  $x \geq 0$  and 0 if  $x < 0$ .

- $E(Q_k) = \int_0^\infty x f(x; k) dx = k$

- $Var(Q_k) = E(Q_k^2) - E(Q_k)^2 = \int_0^\infty x^2 f(x; k) dx - k^2 = 2k$

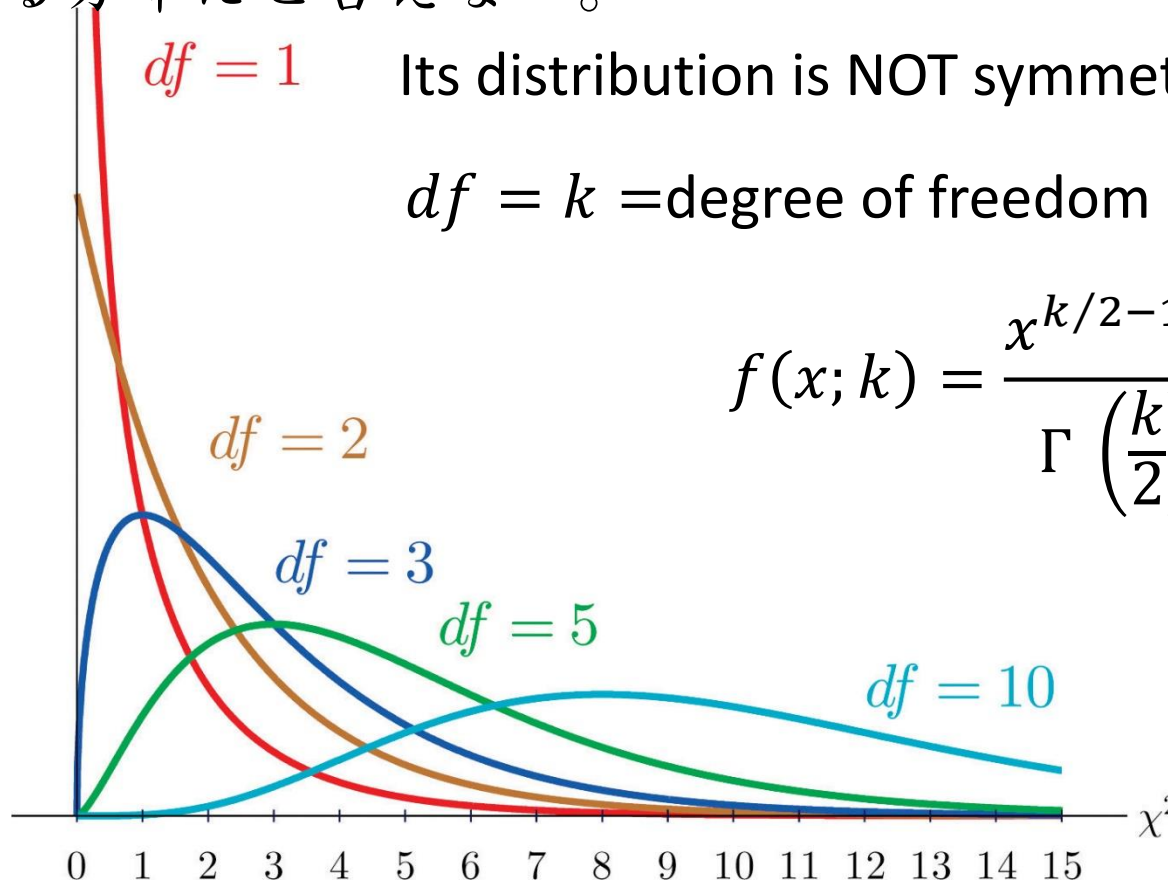
# Some pdf plots of the chi2-distrib

- Used **a lot** in statistics but does not describe natural phenomena like the normal, or exponential distributions.  
統計学においてよく使われているが、自然な現象を記述する分布だと言えない。

- $df = 1$  Its distribution is NOT symmetric

$df = k$  =degree of freedom 自由度

$$f(x; k) = \frac{x^{k/2-1} e^{-x/2}}{\Gamma\left(\frac{k}{2}\right) 2^{k/2}}$$



# Sample variance and $\chi^2$ distribution

- Sample  $X_1, \dots, X_n$  i.i.d.r.v. from a normally distributed population  $N(\mu, \sigma^2)$ .
- Unbiased 不偏 sample variance is ([Lecture 6 slide 10](#))  
$$s_n^2 = \frac{n}{n-1} \bar{\sigma}_X^2 = \frac{1}{n-1} \left( (X_1 - \bar{X})^2 + \dots + (X_n - \bar{X})^2 \right)$$
- $E(s_n^2) = \sigma^2$  (*over sampling distrib. 標本分布上に*)

**Theorem 2** (Helmert/Cochran): The (unbiased) sample variance  $s_n^2$  follows a  $\chi_{n-1}^2$  distribution.

$$(n-1) \frac{s_n^2}{\sigma^2} \sim \chi_{n-1}^2$$

# Approximation (近似) of $\sigma^2$ by $s_n^2$ : one sample is enough ?

$$\blacktriangleright \text{Var}_{\text{samp.distr.}}(s_n^2) = \frac{\sigma^4}{(n-1)^2} \text{Var}(\chi_{n-1}^2) = \frac{2\sigma^4}{n-1}$$

- **When  $n$  is large** the variance of the sample variance  $s_n^2$  is very small (=samples whose sample variance are far away from the population variance  $\sigma$  are rare).  
標本大きさ  $n$  が増加するにつれて、標本分散  $s_n^2$  の分散は小さくなる (=母分散  $\sigma^2$  から離れている標本分散のある標本が珍しい)
- Therefore, **when  $n$  is large** the sample variance is **almost surely** a good approximation of the variance of the population. 標本大きさ  $n$  が大きいときに、標本分散は母集団の分散のよい近時を与えることは **ほとんど確実に** ある。

# Chi2-test for variance (small sample)

- **Assumption:** sample  $X_1, \dots, X_n$  from a  $N(\mu, \sigma^2)$  normally distributed population.
- Don't care about  $\mu$ , want to **have some idea about  $\sigma$** .
- **Null hypothesis  $H_0$ :**  $\sigma = \sigma_0$  ( $\sigma_0$  fixed by you)
- **Test statistic:**  $x^2 = \frac{(n-1)s^2}{\sigma_0^2}$ ,  $f(x^2 | H_0)$  is the pdf of  $\chi_{n-1}^2$
- **P-values:**
  - (i)  $p_r = P(X^2 > x^2)$  (right-sided:  $H_A: \sigma > \sigma_0$ )
  - (ii)  $p_l = P(X^2 < x^2)$  (left-sided:  $H_A: \sigma < \sigma_0$ )
  - (iii)  $2 \min\{p_r, p_l\} = 2 \min\{p_r, 1 - p_r\}$  (2-sided:  $H_A: \sigma \neq \sigma_0$ )
- **Reject  $H_0$  if:**
  - (i)  $p_r \leq \alpha$
  - (ii)  $p_l \leq \alpha$
  - (iii)  $p_r \leq \alpha/2$  or  $p_l \leq \alpha/2$  (same as  $2 \min\{p_r, 1 - p_r\} \leq \alpha$ )

# Practice example (lightbulbs = 白熱電球)

- Lightbulbs manufacturer wants to estimate the lifetime (in hours) of their product.
  - Assumption: lifetime follows  $N(\mu, \sigma^2)$ .
  - They measured the lifetime of 5 lightbulbs and got  $x_1 = 983, x_2 = 1063, x_3 = 1241, x_4 = 1040, x_5 = 1103$
1. Compute the sample mean  $\bar{x}$  and (unbiased) sample variance  $s^2$
  2. Test the assumption  $H_0: \sigma^2 = 4000$  with a one-sided and 2-sided chi2-test.

Answer:  $\bar{x} = 1086, s^2 = \frac{1}{4}(103^2 + 23^2 + 155^2 + 46^2 + 17^2) = 37568.$

$\chi^2 = \frac{4s^2}{4000} = 37.568$  ( $P(X^2 > \chi^2) \approx 0.00$ ). Reject if 1-sided, reject if 2-sided at significance level 0.001.



# Chapter 4: NHST

## Section 4.3: the Student t-test

### Back to the Z-test and large sample

- Review of z-test learned in Lecture 7, [slide 11](#).
  - Need the population mean to be normally distributed  
母集団の対象パラメータは正規分布 $N(\mu, \sigma^2)$ に従う。
  - Need to know the population variance  
母分散 $\sigma^2$ が既存だと必要である。
- If the **sample is large**, no need of these assumptions:  
もしも標本サイズが十分大きいと、上記の仮定がだいぶ不要になる。
- Large-sample usually means **at least**  $n \geq 30 \sim 50$ 
  - Why? Because then the CLT applies 中心極限定理は有効になるからである

# Z-test for the mean: Large-sample (unknown distribution and variance)

- CLT:  $\bar{X} \approx N\left(\mu, \frac{\sigma^2}{n}\right)$  (no need to assume that the population is normally distributed) 対象となる母数の分布は正規分布に従うと仮定しなくてもいい).
- Moreover the sample variance  $s^2$  is a good approximation of the population variance  $\sigma^2$ . さらに、標本分散  $s^2$  は母分散のよい近似を与える。
  - We do not need to know  $\sigma$  and can use instead  $s$ .
  - 母標準偏差  $\sigma$  が未知でも構わない、その代わりに標本偏差  $s^2$  を使うことができる。

➤ Possible to use  $z = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$  instead of  $z = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$  to estimate the mean even for non-normally distributed population

# Small sample: t-test

## 1) Student's $t$ distribution



William Gosset

- $Z = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \sim N(0,1)$
- $Q_n = (n - 1) \frac{S^2}{\sigma^2} \sim \chi_{n-1}^2$  (see Theorem 2 [slide 4](#))

- **Theorem 3** (Gosset = Student, Fisher)

$$T = \frac{Z}{\sqrt{Q_n/n - 1}} = \frac{\bar{X}_n - \mu}{S/\sqrt{n}} \sim t_{n-1}$$

- Where  $t_{n-1}$  is the Student's  $t$ -distribution.
  - $n - 1$  is the degree of freedom (df) = 自由度
- **Remark:** the unknown variance  $\sigma^2$  has been canceled out !

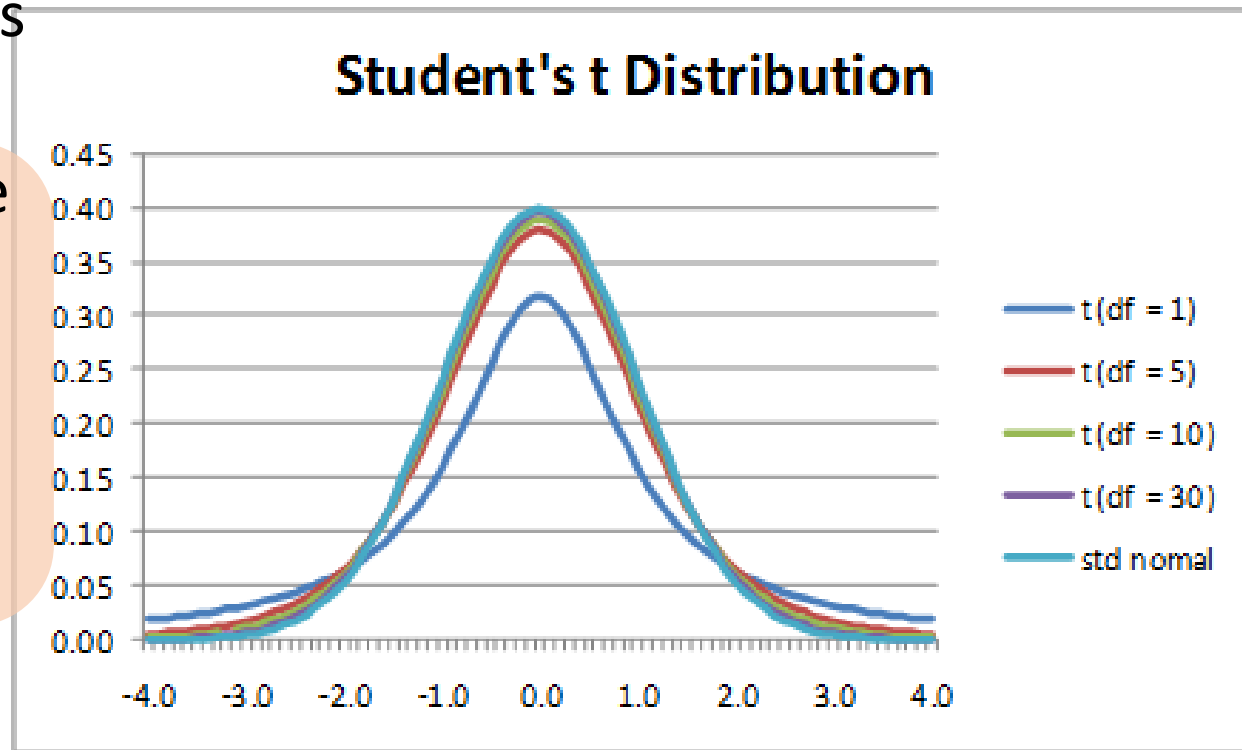
- The pdf is :  $f(x; n - 1) = \frac{\Gamma(\frac{n}{2})}{\Gamma(\frac{n-1}{2})} \left(1 + \frac{x^2}{n-1}\right)^{-n/2}$

Don't  
care

# Small sample: t-test. 2) T-distribution (II)

$$f(x; n - 1) = \frac{\Gamma\left(\frac{n}{2}\right)}{\Gamma\left(\frac{n-1}{2}\right)} \left(1 + \frac{x^2}{n-1}\right)^{-n/2}, \quad \lim_{n \rightarrow \infty} f(x; n) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

- The distribution is symmetric.
- When the degree of freedom  $df$  increases, the distribution approaches the  $N(0,1)$ .
- $E(T_n) = 0$
- $Var(T_n) = \frac{n}{n-2}$



# 1-sample $t$ -test for the mean (unknown variance, small sample)

- **Data:** we assume normal data with both  $\mu$  and  $\sigma$  **unknown:**  
 $x_1, x_2, \dots, x_n \sim N(\mu, \sigma^2)$ .

- **Null hypothesis  $H_0$ :**  $\mu = \mu_0$  for some specific value  $\mu_0$ .

- **Test statistic:**

$$t = \frac{\bar{x}_n - \mu_0}{s/\sqrt{n}} \text{ (Studentized mean)}$$

where  $s^2 = \frac{1}{n-1} \left( (x_1 - \bar{x})^2 + \dots + (x_n - \bar{x})^2 \right)$  is the (unbiased) sample variance.

- **Null distribution:**  $f(t | H_0)$  is the pdf of  $T \sim t_{n-1}$ , the  $t$  distribution with  $n - 1$  degrees of freedom.

- **p-value:**  
two-sided  $p = P(|T| > |t|)$   
Left-sided:  $p = P(T < t)$   
Right-sided:  $p = P(T > t)$

# *t*-test works for small sample

- When the sample is small:
  - The CLT does not hold and we cannot say that  $\bar{X} \approx N\left(\mu, \frac{\sigma^2}{n}\right)$   
中心極限定理が適用されない。
  - ✓ However, if the population is (approximately) normally distributed then the r.v. sample mean  $\bar{X}$  is (approx.) normally distributed. 標本平均はほぼ正規分布で近似できる。
  - But, the sample variance  $S^2$  may not be a good approximation of the population variance  $\sigma^2$ .  
しかし、(不偏の)標本分散 $S^2$ は母分散 $\sigma^2$ からよい近似できないかもしれない。
  - ✓ Better use a *t*-test ☞ no population variance required in the *t*-statistic. ここで母分散が使われていない。
- **Question:** If we don't know  $\sigma$ , why not using a chi2-test (page 8) to estimate it?
- Answer: test says whether  $\sigma \neq \sigma_0$  for some estimated value  $\sigma_0$ . Not that  $\sigma = \sigma_0$  (unless we can compute the power). Population must be very close to be normally distributed.

# Practice: $z$ and 1-sample $t$ -test

- For both questions use significance level  $\alpha = .05$ .
  - Assume the data 2, 4, 4, 10 is drawn from a  $N(\mu, \sigma^2)$ .
  - $H_0: \mu = 0$                        $H_A: \mu \neq 0$
1. Assume  $\sigma^2 = 16$  is known and test  $H_0$  against  $H_A$ .
  2. Now assume  $\sigma^2$  is unknown and test  $H_0$  against  $H_A$

# Chapter 4: NHST

## Section 4.4 Two-sample t-test

POPULATION 1 母集団

**Parameters (母数):**

Mean 母平均:  $\mu_1$

Variance 母分散:  $\sigma_1^2$

.....

SAMPLE 1  
標本

POPULATION 2

**Parameters**

Mean 母平均:  $\mu_2$

Variance 母分散:  $\sigma_2^2$

.....

SAMPLE 2  
標本

Compare  $\mu_1$  and  $\mu_2$   
(when  $\sigma_1 = \sigma_2$ )

SAMPLE 1

**STATISTICS (統計量)**

Size:  $n_1$

mean:  $\bar{x}_1$

Variance:  $s_1^2$

..

Infer 推測

Compare  $\bar{x}_1$  and  $\bar{x}_2$

SAMPLE 2

**STATISTICS (統計量)**

Size:  $n_2$

mean:  $\bar{x}_2$

Variance:  $s_2^2$

..

Independent samples  $P(x \in \Omega_1, y \in \Omega_2) = P(x \in \Omega_1)P(y \in \Omega_2)$



# Two-sample t-test ( $\sigma_1 = \sigma_2$ ) in practice

- **Data:** we assume **normal data** with  $\mu_1, \mu_2$  and **(same)  $\sigma$**  unknown:

$$x_1, \dots, x_{n_1} \sim N(\mu_1, \sigma^2), \quad y_1, \dots, y_{n_2} \sim N(\mu_2, \sigma^2)$$

- **Null hypothesis  $H_0$ :**  $\mu_1 = \mu_2$ .

- **Pooled variance:**  $s_p^2 = \frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1+n_2-2}$

where  $s_1^2$  (resp.  $s_2^2$ ) is the sample variance of the  $x$  (resp  $y$ ) sample.

- **Test statistic:**

$$t = \frac{\bar{x} - \bar{y}}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

- **Null distribution:**  $f(t | H_0)$  is the pdf of  $T_{n_1+n_2-2} \sim t_{n_1+n_2-2}$

- **P-value:**  $P(T_{n_1+n_2-2} > t)$  (right-sided,  $H_A: \mu_1 > \mu_2$ )  
 $P(T_{n_1+n_2-2} < t)$  (left-sided,  $H_A: \mu_1 < \mu_2$ )  
 $P(|T_{n_1+n_2-2}| > |t|)$  (two-sided,  $H_A: \mu_1 \neq \mu_2$ )

# Practice of the 2-sample $t$ -test for the mean

Real data from 1408 women admitted to a maternity hospital for

実のデータによると、産科病院に入院した1408人の女の人を二つの理由で以下のように分かれている：

(i) **medical** (booked) reasons or through  
医療上の（予約）理由で

または

(ii) unbooked **emergency** admission.  
（予約されない）急患診療。

The duration of pregnancy is measured in complete weeks from the beginning of the last menstrual period.

最後の月経期から週数によって妊娠期間が計られる。

- (i) **Medical**: 775 obs. with  $\bar{x} = 39.08$  and  $s^2 = 7.77$ .
  - (ii) **Emergency**: 633 obs. with  $\bar{x} = 39.60$  and  $s^2 = 4.95$
1. Set up and run a two-sample t-test to investigate whether the duration differs for the two groups.
  2. What assumptions did you make?

# Some Remarks about 2-sample t-test

- Possible to use two-sample Z-test as well if:
  - Sample sizes  $n_1$  and  $n_2$  are large  $\geq 30 \sim 50$  and the populations variances  $\sigma_1^2$  and  $\sigma_2^2$  are known  
標本数  $n_1$  と  $n_2$  がともに 30 ~ 50 以上であり、母集団分散  $\sigma_1^2$  と  $\sigma_2^2$  が既存であるとき
  - Or both population follows a normal distribution and  $\sigma_1^2$  and  $\sigma_2^2$  are known 二つの母集団の分布が正規分布に従い、 $\sigma_1^2$  と  $\sigma_2^2$  が既存であるとき

$$Z = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{\sigma_1^2/n_1 + \sigma_2^2/n_2}}$$

- **Welch's test:** case where  $\sigma_1 \neq \sigma_2$  (未知、unknown)

- Then  $t = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{s_1^2/n_1 + s_2^2/n_2}}$  follows  $t$ -distribution

- Degree of freedom !  $df = \left\lfloor \frac{(s_1^2/n_1 + s_2^2/n_2)^2}{(s_1^2/n_1)^2/(n_1-1) + (s_2^2/n_2)^2/(n_2-1)} \right\rfloor$

# Control (placebo) group vs treatment group Example “The bad scientist...”

- Suppose that a scientist Jerry is terrible at designing effective treatment but he always carefully randomly divides his patients into control and treatment groups.  
科学者Jerryは効き目のある治療を設計するのが下手だが、常に患者を独立無作為にコントロール（対照群）と治療群の2つに分ける。
- His null hypothesis  $H_0$  is that the treatment is no better than the placebo.  
帰無仮説は治療はプラセボより効くことはない。
- He uses a significance level of  $\alpha = 0.05$ .
- If his p-value is less than  $\alpha$  he publishes a paper claiming the treatment is significantly better than a placebo.  $p$ 値が0.05の有意水準より小さいと、治療はプラセボよりも有意に効くと判断し、論文を投稿し出版する。

1. Since his treatments are never, in fact, effective what percentage of his experiments result in published papers?

Jerryが行った実験が科学論文になった割合を計算するにはどうすればよいか。

*Hint: Think about significance level and type of error.*

2. What percentage of his published papers describe treatments that are better than placebo?

実際に効能のある治療を記述する科学論文の割合はどのくらいか。

Answer: 1. 5%.  $\alpha = P(\text{type I error}) = P(\text{reject a correct } H_0) = 5\%$

2. None.

## ..and the good scientist”

Jenna is a genius at designing treatments, all her proposed treatments are effective. She always tests her new treatment with control/treatment group,

Jennaは治療の設計に対して才能があり、提案した治療はすべて効果がある。帰無仮説を彼女はコントロール（対照群）と治療群に対して新しくテストし、

帰無仮説を

and set the null hypothesis to be

“ $H_0$ : the treatment is not more effective than placebo”

「治療はプラセボより効くことはない」

and runs a two-sample t-test at significance level  $\alpha = 0.05$ .

をとし、0.05有意水準 2つの標本t-検定を行う。

She publishes a paper if her p-value  $< \alpha$ .

p値  $< 0.05$  成り立ったら論文を投稿し出版する。

1. How could you determine what percentage of her experiments result in publications?

Jennaが行った実験が科学論文になった割合を計算するにはどうすればよいか。

Hint: Think about significance level and type of error.

2. What percentage of her published papers describe effective treatments?

実際に効能のある治療を記述する科学論文の割合はどのくらいか。

Answer: 1.  $1 - p(\text{type II error}) = \text{power of the treatment.}$

If a tiny better than placebo, then nearly 5%. If much better than placebo then virtually 100%.

2. All.



# Chapter 4: NHST

## Section 4.5 Paired difference sampling

### 対応のあるデータ

- **Example:** Test a teaching method on slow-learners kids.  
学業遅滞児に対する指導法を検定する。
- reading IQs: measure ability to learn to read for kids ( $\mu = 100$ ). 子供に関して、読書能力を獲得できることをはかるもの：読書IQ
- Let 8 pairs ( $n=16$ ) of slow-learners kids.  
学業遅滞児の8ペア(16子)を考える。
- The two kids in each pair have similar reading IQs.  
各ペアにいる二人の子は同等な読書IQがある。
- For each pair すべて子のペアに対して
  - A kid is randomly selected to learn a new method (N)  
新指導法を学ぶ子を無作為に選ぶ。
  - The other kid learns the standard method (S)  
他の子は標準法を学ぶ。

# Example: Test results

Pair Nb	New (N) Method	Standard Method (S)
1	77	72
2	74	68
3	82	76
4	73	68
5	87	84
6	69	68
7	66	61
8	80	76

- 2-sample t-test:

- $\bar{x}_N = 76, \quad s_N = 6.93$

- $\bar{x}_S = 71.63, \quad s_S = 7.01$

- Test:  $H_0$ : New method is not better than Standard.  $\mu_N = \mu_S$   
 $H_A$ : New method is better than the Standard.  $\mu_N > \mu_S$

- Pooled sample:  $s_P^2 = \frac{(n_1-1)s_N^2 + (n_2-1)s_S^2}{n_1+n_2-2} = \frac{7 \cdot 6.93^2 + 7 \cdot 7.01^2}{8+8-2}$   
 $s_P^2 = 48.55, s_P \approx 6.9687,$

- $$t = \frac{\bar{x}_N - \bar{x}_S}{s_P \cdot \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} = \frac{4.37}{3.4823} \approx 1.26$$

- T-table (df=14): 1-sided p-value  $> 0.1$   
 ─ don't reject at 0.1 significance level

## Example (II): something is wrong ?

- However, the data clearly suggests that the new method is better than the standard one.  
ただし、データによると新方法は標準法よりも効果があるとはっきり推定できる。
- What's wrong ?  
何かおかしい？
  - The samples on the two groups are not independent!!  
二つのグループの標本が独立ではない
  - The two-sample t-test is not valid, the pooled variance  $s_p^2$  is large compared to the small sample means difference  $\bar{x}_N - \bar{x}_S$ .  
合併分散  $s_p^2$  は標本平均の差  $\bar{x}_N - \bar{x}_S$  に比して大きすぎ、妥当ではない。
    - The test cannot detect such small differences.  
検定はこのような小さな差を検出できない
- What can we do?
  - Use a 1-sample t-test on  $\mu_D = \mu_N - \mu_S$ .

## Example (III): paired differences 一对比較

Pair Nb	New Method	Standard Method	Difference (N-S)
1	77	72	5
2	74	68	6
3	82	76	6
4	73	68	5
5	87	84	3
6	69	68	1
7	66	61	5
8	80	76	4

- $H_0: \mu_D = 0$  ( $\mu_N - \mu_S = 0$ )       $H_A: \mu_D > 0$  ( $\mu_N > \mu_S$ )
- 1-sample t-test:  $t = \frac{\bar{x}_D - 0}{s_D \sqrt{n_D}} = \frac{4.375}{1.685/\sqrt{8}} = 7.34$
- T-table (df= $n_D - 1=7$ ) gives: p-value  $< 0.0005$  **reject  $H_0$**

# Paired difference NHST 一对比較法

- $H_0: \mu_D = 0$  (or any constant  $D_0$ )
- $H_A: \mu_D \neq 0$  or  $< 0$  or  $> 0$  (or  $\neq$  or  $>$  or  $< D_0$ )
- **Large sample (z-test)**
  - Test statistic:  $z = \frac{\bar{x}_D - D_0}{s_D / \sqrt{n_D}} \approx \frac{\bar{x}_D - D_0}{S_D / \sqrt{n_D}}$  (follows  $N(0,1)$  under  $H_0$ )
  - P-value:  $P(Z > z)$  or  $P(|Z| > |z|)$  or  $P(Z < z)$

## Assumption

- Sample size  $n_D$  is large (usually **at least** 30~50)
- **Small sample (t-test)**
  - Test statistic:  $t = \frac{\bar{x}_D - D_0}{s_D / \sqrt{n_D}}$  (follows  $t_{n_D-1}$ )
  - P-value:  $P(T_{n_D-1} > t)$  or  $P(|T_{n_D-1}| > |t|)$  or  $P(T_{n_D-1} > t)$

## Assumption

- population of difference is (approx.)  $\sim N(\mu_D, \sigma_D^2)$

# Chapter 4: NHST

## Section 4.6 Comparing two population variances: F-test

POPULATION 1 母集団

Mean 母平均:  $\mu_1$

Variance 母分散:  $\sigma_1^2$

Assumption:  $N(\mu_1, \sigma_1^2)$

SAMPLE 1  
標本

POPULATION 2

Mean 母平均:  $\mu_2$

Variance 母分散:  $\sigma_2^2$

Assumption:  $N(\mu_2, \sigma_2^2)$

SAMPLE 2  
標本

Compare  $\sigma_1$  and  $\sigma_2$   
(Both population are normally distributed)

SAMPLE 1  
STATISTICS (統計量)

Size:  $n_1$

mean:  $\bar{x}_1$

Variance:  $s_1^2$

..

Infer 推測

Check if  $s_1^2/s_2^2 \approx 1$

SAMPLE 2  
STATISTICS (統計量)

Size:  $n_2$

mean:  $\bar{x}_2$

Variance:  $s_2^2$

..

Independent samples  $P(x \in \Omega_1, y \in \Omega_2) = P(x \in \Omega_1)P(y \in \Omega_2)$

# The F-distribution

- **Theorem (Fisher-Snedecor)**

If  $X \sim \chi_{n_1}^2$  and  $Y \sim \chi_{n_2}^2$  then the random variable  $\frac{X/n_1}{Y/n_2}$  has a distribution called  $F$  with  $n_1$  and  $n_2$  degrees of freedom.

- The density function is complicated:

$$f(x; n_1, n_2) = \frac{(n_1/n_2)^{n_1/2} x^{n_1-1/2}}{B(n_1/2, n_2/2)} \left(1 + \frac{n_1}{n_2} x\right)^{-n_1-n_2/2}$$

where  $B(n_1/2, n_2/2) = \int_0^1 t^{n_1-2/2} (1-t)^{n_2-2/2} dt$

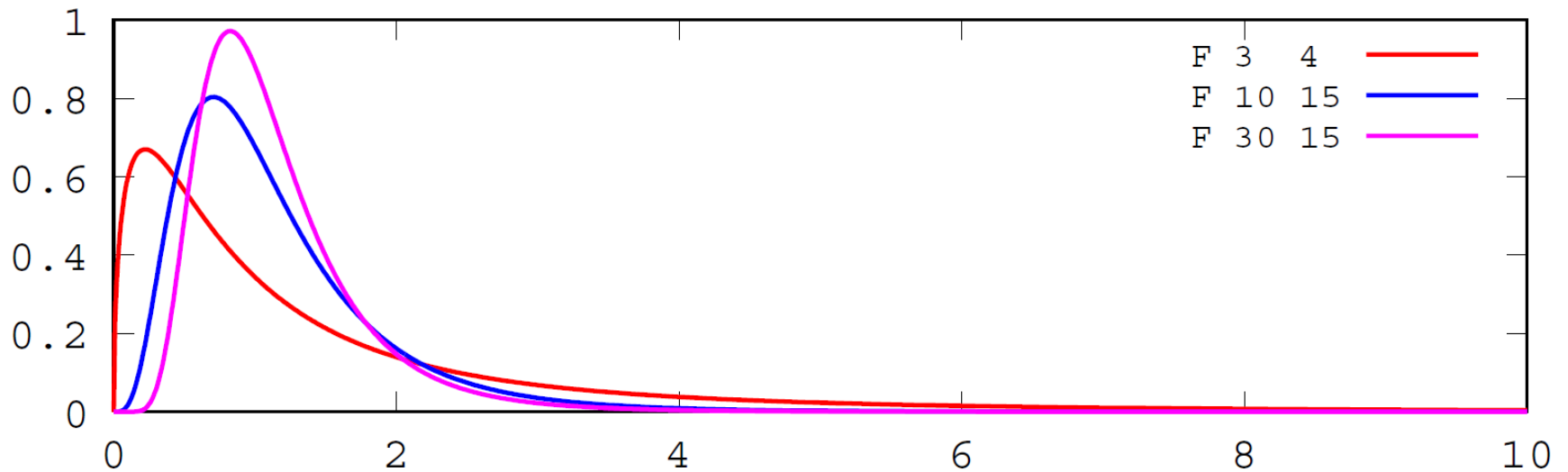
- If  $F \sim F_{n_1, n_2}$  then  $E(F) = \frac{n_2}{n_2-2}$  and

- $Var(F) = \frac{2n_2^2(n_1+n_2-2)}{n_1(n_2-2)^2(n_2-4)}$

# F-distribution plot

- Not symmetric 対称ではない.
- Range  $[0; \infty)$  領域

Plot of F distributions





# The F-test of equality of variance

- Recall the Theorem 1 (page [7](#)) and Fisher-Snedecor's theorem (page [30](#))

- $X = (n_1 - 1) \frac{s_1^2}{\sigma_1^2} \sim \chi_{n_1-1}^2$  and  $Y = (n_2 - 1) \frac{s_2^2}{\sigma_2^2} \sim \chi_{n_2-1}^2$

- $\triangleright \frac{X/n_1-1}{Y/n_2-1} = \frac{s_1^2/\sigma_1^2}{s_2^2/\sigma_2^2} \sim F_{n_1-1, n_2-1}$

- Null hypothesis  $H_0$ :**  $\sigma_1 = \sigma_2$

- Test statistic:** Under  $H_0$   $\frac{s_1^2/\sigma_1^2}{s_2^2/\sigma_2^2} = \frac{s_1^2}{s_2^2}$  so  $f = \frac{s_1^2}{s_2^2} \sim F_{n_1-1, n_2-1}$

- p-value:**  $P(F_{n_1-1, n_2-1} > f)$  (2-sided  $H_A: \sigma_1 \neq \sigma_2$ )  
(right-sided:  $H_A: \sigma_1^2 > \sigma_2^2$ , NOT  $\sigma_2^2 > \sigma_1^2$ )

- Reject if  $p < \alpha$  (1-sided) or  $p < \alpha/2$  (2-sided).

- Assumption:**

- Both populations are normally distributed, NOT only approximately.

- So this test is not used a lot in practice** あまり使われていない

# Example I

- Scores at a Mathematic MCQ test shows (source: 1998, *American Educational Research J.*)

	Males	Females
Sample size	1,764	1,739
Mean	48.4	48.9
Standard Deviation	12.96	11.85

- 数学の多項  
選択試験で  
得られた点数 :

- Test the hypothesis that the variance of the male's score is more variable than the females.  
男の点数の分散の方が女より変わりやすいという仮説を検定せよ。

- $H_0: \sigma_M = \sigma_F$   $f = s_M^2 / s_F^2 = 1.19612$
- 1-sided test using  $F_{\infty, \infty}$ -table:  $\alpha = 0.01$  level gives  $c_\alpha = 1 < f$  so we reject  $H_0$  in favor of  $\sigma_M > \sigma_F$
- (Using computer we find that  $F_{1763, 1738}(f) = .999909$ )

# Practice problem: back to slow learners

- Remember exercise page [26](#) on slow-learner kids.
  - The t-test was made under the assumption that  $\sigma_N = \sigma_S$ .
1. Test this hypothesis at significance level  $\alpha = 0.1$
  2. Can we conclude that  $\sigma_N = \sigma_S$  ?